## Distributionally Robust Batch Contextual Bandits

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- Motivation: distributional shifts in batch contextual bandit
- Distributionally robust formulation
- Oistributionally robust policy learning
- Numerical results
- 5 Extension to *f*-divergence uncertainty set

### Motivation: distributional shifts in batch bandit





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## Motivation: distributional shifts in batch bandit





A collection of triplets of context, action and rewards in an environment  $\mathbf{P}_{a}$ .





We aim to deploy a robust policy in unknown environments  $\mathbf{P}_{\rm b}$  which are similar but slightly different from the previous environment.

 $\bm{P}_{\rm b} \approx \bm{P}_{\rm a}$ 

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# Main challenges



• Incomplete (bandit-type) data:

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• Distributional shifts: covariate shift and concept drift.



# Setting



• Context:  $X \in \mathcal{X}$ ; Actions:  $A \in \mathcal{A} = \{a^1, a^2, \dots, a^d\}$ ; Rewards:  $(Y(a^1), Y(a^2), \dots, Y(a^d)) \in \prod_{j=1}^d \mathcal{Y}_j$ .

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- Batch bandit data:  $\{(X_i, A_i, Y_i(A_i))\}_{i=1}^n$ , where  $(X_i, Y_i(a^1), Y_i(a^2), \dots, Y_i(a^d)) \stackrel{i.i.d.}{\sim} \mathbf{P}_0$ , and  $A_i \sim \pi_0(\cdot \mid X_i)$  is known.

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- Goal: learn a robust policy that performs well in the presence of unknown distributional shifts.

# Assumptions: standard assumptions<sup>1</sup>



### Assumption (Standard assumptions)

 Unconfoundedness: (Y(a<sup>1</sup>), Y(a<sup>2</sup>),..., Y(a<sup>d</sup>)) is independent with A conditional on X, i.e.,

 $(Y(a^1), Y(a^2), \ldots, Y(a^d)) \perp A \mid X.$ 

- 2. Overlap: There exists some  $\eta > 0$ ,  $\pi_0(a \mid x) \ge \eta$ ,  $\forall (x, a) \in \mathcal{X} \times \mathcal{A}$ .
- 3. Bounded reward support:  $0 \le Y(a^i) \le M$  for i = 1, 2, ..., d.

<sup>1</sup>This assumption is standard and commonly adopted in both the causal inference literature (Rosenbaum and Rubin [1983], Imbens [2004], Imbens and Rubin [2015]) and the policy learning literature (Zhang et al. [2012], Zhao et al. [2012], Kitagawa and Tetenov [2018], Swaminathan and Joachims [2015], Zhou et al. [2017]).

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# Assumptions: positive densities/probabilities

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### Assumption (Positive densities/probabilities)

- 1. Continuous case: for any i = 1, 2, ..., d,  $Y(a^i)|X$  has a conditional density  $f_i(y_i|x)$ , and  $f_i(y_i|x) \ge \underline{b} > 0$  over the interval [0, M] for any  $x \in \mathcal{X}$ .
- 2. Discrete case: for any i = 1, 2, ..., d,  $Y(a^i)$  supported on a finite set  $\mathbb{D}$ , and  $P_0(Y(a^i) = v|X) \ge \underline{b} > 0$  for any  $v \in \mathbb{D}$ .





• How to model distributional shifts?

• Kullback-Leibler divergence:  $KL(\mathbf{P}||\mathbf{P}_0) \triangleq \int_{\mathcal{X} \times \prod_{j=1}^d \mathcal{Y}_j} \log\left(\frac{\mathrm{d}\mathbf{P}}{\mathrm{d}\mathbf{P}_0}\right) \mathrm{d}\mathbf{P}.$ 



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- Uncertainty set:  $\mathcal{U}_{\mathbf{P}_0}(\delta) = \{\mathbf{P} \mid \mathcal{K}\mathcal{L}(\mathbf{P}||\mathbf{P}_0) \le \delta\}.$



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- Uncertainty set:  $\mathcal{U}_{\mathbf{P}_0}(\delta) = \{\mathbf{P} \mid KL(\mathbf{P}||\mathbf{P}_0) \le \delta\}.$
- Distributionally robust value function (population level):

$$Q_{\mathrm{DRO}}(\pi) \triangleq \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_0(\delta)}} \mathbf{E}_{\mathbf{P}}[Y(\pi(X))].$$



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- Distributionally robust value function (population level):

$$Q_{\text{DRO}}(\pi) \triangleq \inf_{\substack{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_0(\delta)} \\ \text{Infinite dimensional optimization.} \\ \text{Bandit observations for } \mathbf{P}_0.} \mathbf{E}_{\mathbf{P}}[Y(\pi(X))].$$

• Strong duality<sup>2</sup> for the distributionally robust value function:

$$Q_{\text{DRO}}(\pi) = \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_{0}(\delta)}} \mathbf{E}_{\mathbf{P}}[Y(\pi(X))]$$
  
= 
$$\sup_{\alpha \ge 0} \{-\alpha \log \mathbf{E}_{\mathbf{P}_{0}} \left[\exp(-Y(\pi(X))/\alpha)\right] - \alpha \delta\}$$



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where  $\mathbf{P}_0 * \pi_0$  denotes the product distribution on the space  $\mathcal{X} \times \prod_{i=1}^d \mathcal{Y}_i \times \mathcal{A}$ .

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where  $\mathbf{P}_0 * \pi_0$  denotes the product distribution on the space  $\mathcal{X} \times \prod_{j=1}^d \mathcal{Y}_j \times \mathcal{A}$ . • Finite-sample estimate:  $\hat{Q}_{\text{DRO}}(\pi) = \sup_{\alpha \ge 0} \{-\alpha \log \hat{W}_n(\pi, \alpha) - \alpha \delta\}$ , where

$$\hat{W}_n(\pi,\alpha) = \frac{1}{n} \sum_{i=1}^n \frac{\exp(-Y_i(A_i)/\alpha) \mathbf{1}\{\pi(X_i) = A_i\}}{\pi_0(A_i \mid X_i)}$$

<sup>2</sup>Hu and Hong [2013]

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$$\hat{W}_{n}(\pi,\alpha) = \underbrace{\frac{1}{\sum_{i=1}^{n} \frac{\mathbf{1}\{\pi(X_{i}) = A_{i}\}}{\pi_{0}(A_{i}|X_{i})}}_{\text{More stable}} \sum_{i=1}^{n} \frac{\exp(-Y_{i}(A_{i})/\alpha)\mathbf{1}\{\pi(X_{i}) = A_{i}\}}{\pi_{0}(A_{i} \mid X_{i})}.$$
<sup>2</sup>Hu and Hong [2013]
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# Central limit theorem

#### Theorem

Under assumptions mentioned earlier, for any policy  $\pi \in \Pi$ , we have

$$\sqrt{n}\left(\hat{Q}_{\mathrm{DRO}}(\pi)-Q_{\mathrm{DRO}}(\pi)\right)\Rightarrow\mathcal{N}\left(0,\sigma^{2}(\alpha^{*})\right),$$

where  $\alpha^*$  is the optimal dual variable, defined by

$$\alpha^* = \underset{\alpha \ge 0}{\arg \max} \left\{ -\alpha \log \mathbf{E}_{\mathbf{P}_0} \left[ \exp(-Y(\pi(X))/\alpha) \right] - \alpha \delta \right\},\$$

and  $\sigma^2(\alpha) =$ 

$$\frac{\alpha^2}{\mathsf{E}[\exp\left(-Y(\pi(X))/\alpha\right)]^2} \mathsf{E}\left[\frac{1}{\pi_0\left(\pi(X)|X\right)}\left(\exp\left(-Y(\pi(X))/\alpha\right) - \mathsf{E}\left[\exp\left(-Y(\pi(X))/\alpha\right)\right]\right)^2\right]$$





• How to find a good policy:  $\arg \max_{\pi \in \Pi} Q_{DRO}(\pi)$ ?



- How to find a good policy:  $\arg \max_{\pi \in \Pi} Q_{DRO}(\pi)$ ?
- Given a policy class  $\Pi$ , learn a distributionally robust policy:

$$\hat{\pi}_{\text{DRO}} = \arg \max_{\pi \in \Pi} \hat{Q}_{\text{DRO}}(\pi)$$

$$= \arg \max_{\pi \in \Pi} \sup_{\alpha \ge 0} \{-\alpha \log \hat{W}_n(\pi, \alpha) - \alpha \delta\}$$



- How to find a good policy:  $\arg \max_{\pi \in \Pi} Q_{DRO}(\pi)$ ?
- Given a policy class  $\Pi$ , learn a distributionally robust policy:

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- Alternatively update  $\pi$  and  $\alpha$ ;
  - Using Newton-Raphson method to update  $\alpha$ ; converge fast empirically.



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- Given a policy class II, learn a distributionally robust policy:

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- Alternatively update  $\pi$  and  $\alpha$ ;
  - Using Newton-Raphson method to update  $\alpha$ ; converge fast empirically.
- How does  $\hat{\pi}_{\mathrm{DRO}}$  perform?

$$R_{\text{DRO}}(\hat{\pi}_{\text{DRO}}) = \max_{\pi' \in \Pi} Q_{\text{DRO}}(\pi') - Q_{\text{DRO}}(\hat{\pi}_{\text{DRO}})$$
  
= 
$$\max_{\pi' \in \Pi} \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_0}(\delta)} \mathbf{E}_{\mathbf{P}}[Y(\pi'(X))] - \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_0}(\delta)} \mathbf{E}_{\mathbf{P}}[Y(\hat{\pi}_{\text{DRO}}(X))].$$

# Statistical performance guarantee

#### Theorem

Under assumptions mentioned earlier, with probability at least  $1 - \varepsilon$ , we have in the continuous case

$$\mathcal{R}_{ ext{DRO}}(\hat{\pi}_{ ext{DRO}}) \leq rac{4}{\underline{b}\eta \sqrt{n}} \left( (\sqrt{2}+1) \kappa^{(n)}\left(\Pi
ight) + \sqrt{2\log\left(rac{2}{arepsilon}
ight)} + C 
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and in the discrete case

$$R_{ ext{DRO}}(\hat{\pi}_{ ext{DRO}}) \leq rac{4M}{\underline{b}\eta\sqrt{n}} \left( 24(\sqrt{2}+1)\kappa^{(n)}\left(\Pi
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where  $\kappa^{(n)}(\Pi)$  represents the complexity of the policy class  $\Pi$ , and  $\eta > 0$  is a lower bound for the propensity score (collection policy)  $\pi_0(a, x)$  mentioned in the previous assumption.

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# Remarks on the complexity term $\kappa^{(n)}(\Pi)$

#### Example

- Finite class: For a policy class Π<sub>Fin</sub> containing a finite number of policies, we have κ<sup>(n)</sup> (Π<sub>Fin</sub>) ≤ √log(|Π<sub>Fin</sub>|).
- Linear class: For X ⊂ R<sup>p</sup>, each policy π ∈ Π<sub>Lin</sub> is parameterized by a set of d vectors Θ = {θ<sub>a</sub> ∈ R<sup>p</sup> : a ∈ A} ∈ R<sup>p×d</sup>, and the mapping π : X → A is defined as

$$\pi_{\Theta}(x) \in rgmax_{a \in \mathcal{A}} \left\{ heta_{a}^{ op} x 
ight\}.$$

Then, we have  $\kappa^{(n)}(\Pi_{\text{Lin}}) \leq C\sqrt{dp \log(d) \log(dp)}$ .

In general, κ<sup>(n)</sup>(Π) can be bounded by the VC dimension when d = 2, or the graph dimension when d > 2.

# Simulation, real data experiments, and the selection of $\boldsymbol{\delta}$

# Simulation study: benchmark



Benchmark: let  $\overline{\Pi}$  denote the class of all measurable mappings from contexts  $\mathcal{X}$  to the action set  $\mathcal{A}$ .

• Bayes policy  $\overline{\pi}^*$ :

$$\overline{\pi}^* \in rg\max_{\pi \in \overline{\mathsf{\Pi}_0}} \mathsf{E}_{\mathsf{P}}[Y(\pi(X))], ext{ and }$$

• Bayes DRO policy  $\overline{\pi}^*_{\mathrm{DRO}}$ :

$$\overline{\pi}^*_{\mathrm{DRO}} \in \arg\max_{\pi \in \overline{\Pi}} \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_0(\delta)}} \mathbf{E}_{\mathbf{P}}[Y(\pi(X))].$$

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$$\overline{\pi}^*_{\mathrm{DRO}} \in \operatorname*{arg\,max}_{\pi \in \overline{\Pi}} \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_0(\delta)}} \mathbf{E}_{\mathbf{P}}[Y(\pi(X))].$$

- Best policies, but may not in the policy class  $\Pi$ .
- Not learnable, but theoretically easy to compute in the simulation environment, because the policies are the best response for each X.

### Simulation study



• 3 actions; 5-dimensional features, but only the first two matter:

$$Y(i)|X \sim \mathcal{N}(\mu_i(X), \sigma_i^2), ext{ for } i=1,2,3.$$

where the conditional mean  $\mu_i(x)$  and conditional variance  $\sigma_i$  are chosen as

$$\begin{aligned} \mu_1(x) &= 0.2x(1), & \sigma_1 = 0.8, \\ \mu_2(x) &= 1 - \sqrt{(x(1) + 0.5)^2 + (x(2) - 1)^2}, & \sigma_2 = 0.2, \\ \mu_3(x) &= 1 - \sqrt{(x(1) + 0.5)^2 + (x(2) + 1)^2}, & \sigma_3 = 0.4. \end{aligned}$$

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- The linear policy class:  $\Pi = \{\pi(x) = \arg \max_{a \in \mathcal{A}} \ \{\theta_a^\top x\} : \theta_a \in \mathbf{R}^p, a \in \mathcal{A}\}.$

# Non-linear example with the linear policy class

(a) Bayes policy  $\overline{\pi}^*$ ;

(c) Bayes distributionally robust policy  $\overline{\pi}^*_{\mathrm{DRO}}$ 



(b) non-DRO linear policy;

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(d) distributionally robust linear policy  $\hat{\pi}_{DRO}$ .

Figure 1:  $\sigma_1 = 0.8(blue), \sigma_2 = 0.2(orange), \sigma_3 = 0.4(green).$ 

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#### <sup>3</sup>Credit: Getty Images/iStockphoto

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### Backgrounds



• Dataset Description:<sup>4</sup> 180002 data points on whether individuals voted in the 2006 primary election with their characteristics. There is 1 control and 4 treatments.



### Actions

- There are 5 actions (1 control with probability 5/9 and 4 treatments each with probability 1/9).
  - Nothing: No action is performed.
  - **Civic:** A letter with "Do your civic duty" is mailed to the household before the primary election.
  - **Monitored:** A letter with "You are being studied" is mailed to the household before the primary election.
  - **Self History:** A letter with the past voting records of the voter's household is mailed to the household before the primary election.
  - **Neighbors:** A letter with the past voting records of this voter's household and neighbors is mailed to the household.

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- Neighbors is dominant for the whole population. To make all actions comparable, we minus an artificial cost of deploying each action:
   Y<sub>i</sub>(a) = 1{voter i votes in 2006 under action a} c<sub>a</sub>.

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- Neighbors is dominant for the whole population. To make all actions comparable, we minus an artificial cost of deploying each action:
   Y<sub>i</sub>(a) = 1{voter i votes in 2006 under action a} c<sub>a</sub>.
- Goal: learn a distributionally robust policy to maximize voting turnout.

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# Training and evaluation procedure





### Training and evaluation procedure



- We divide the training and test population based on the *city* (101 cities in the dataset).
  - Natural covariate shifts and concept drifts; e.g., the distribution of *year of birth* is generally different across different cities.
  - Leave-one-out to generate 101 pairs of training set and test set.

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|            |                | mean   | std    | min     | 5% quantile |
|------------|----------------|--------|--------|---------|-------------|
| Non-robust |                | 0.0386 | 0.0991 | -0.2844 | -0.1104     |
| Robust     | $\delta = 0.1$ | 0.0458 | 0.0989 | -0.2321 | -0.1007     |
|            | $\delta = 0.2$ | 0.0368 | 0.0895 | -0.2314 | -0.0785     |
|            | $\delta = 0.3$ | 0.0397 | 0.0864 | -0.2313 | -0.0677     |
|            | $\delta = 0.4$ | 0.0383 | 0.0863 | -0.2312 | -0.0677     |

Table 1: Comparison of important statistics for 101 test results.

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non-robust decision tree  $\hat{\pi}_{\text{DRO}}, \delta = 0.1$ 

### Results for $\delta = 0.1$

7

6

Density 5



(a) Comparison of test performances between a distributionally robust decision tree and a non-robust decision tree



#### (b) Example of a distributionally robust tree



### How to select the uncertain size $\delta$ in practice?

Selecting  $\delta$  is more like a managerial decision rather than a scientific procedure.

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- $\bullet\,$  Compute  $\delta\,$  based on the training data:
  - Estimate distributions of Y using any causal inference/machine learning methods.
  - Randomly split training data into 20 cities ( $\mathbf{P}^{20}$ ) against 80 cities ( $\mathbf{P}^{80}$ ) 100 times.
  - Estimate  $\delta$  based on  $KL(\mathbf{P}^{20}||\mathbf{P}^{80}) = KL(\mathbf{P}^{20}_X||\mathbf{P}^{80}_X) + \mathbf{E}_{\mathbf{P}^{20}_Y}[KL(\mathbf{P}^{20}_Y|X||\mathbf{P}^{80}_Y|X)].$



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- Check the performance of  $\hat{\pi}^{\delta}_{\mathrm{DRO}}$  using different value functions.
  - Robust policy does not compromise the non-robust value function.
  - The performance is not sensitive to  $\delta$ , when  $\delta \ge 0.2$ .



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# Extension to *f*-divergence uncertainty set



• Up to now, all of the results are for Kullback-Leibler divergence.

• We can also generalize to  $f_k$ -divergence.

### Extension to *f*-divergence uncertainty set



$$D_k(\mathbf{P}||\mathbf{P}_0) \triangleq \int f_k\left(\frac{d\mathbf{P}}{d\mathbf{P}_0}\right) d\mathbf{P}_0.$$



# Extension to *f*-divergence uncertainty set

For  $f_k(t) \triangleq \frac{t^k - kt + k - 1}{k(k-1)}$ , define *f*-divergence as

$$D_k(\mathbf{P}||\mathbf{P}_0) \triangleq \int f_k\left(\frac{d\mathbf{P}}{d\mathbf{P}_0}\right) d\mathbf{P}_0.$$

#### Theorem

Under assumptions mentioned above, with probability at least  $1 - \varepsilon$ , we have in the continuous case (similar result for the discrete case)

$$\begin{split} & \max_{\pi' \in \Pi} \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_{0}}^{k}(\delta)} \mathbf{E}_{\mathbf{P}}[Y(\pi'(X))] - \inf_{\mathbf{P} \in \mathcal{U}_{\mathbf{P}_{0}}^{k}(\delta)} \mathbf{E}_{\mathbf{P}}[Y(\hat{\pi}_{\mathrm{DRO}}(X)) \\ & \leq \quad \frac{4c_{k}(\delta)}{\underline{b}\eta\sqrt{n}} \left( (\sqrt{2}+1)\kappa^{(n)}(\Pi) + \sqrt{2\log\left(\frac{2}{\varepsilon}\right)} + C \right), \end{split}$$

where  $c_k(\delta) \triangleq (1 + k(k-1)\delta)^{1/k}$ .

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**Si N**, Zhang F, Zhou Z, and Blanchet J. "Distributional Robust Batch Contextual Bandits." arXiv preprint arXiv:2006.05630 (2020). under review.

# Thanks!

#### Reference

### References I

- Alan S Gerber, Donald P Green, and Christopher W Larimer. Social pressure and voter turnout: Evidence from a large-scale field experiment. *American political Science review*, 102(1):33–48, 2008.
- Zhaolin Hu and L Jeff Hong. Kullback-leibler divergence constrained distributionally robust optimization. *Available at Optimization Online*, 2013.
- Guido W Imbens. Nonparametric estimation of average treatment effects under exogeneity: A review. *Review of Economics and statistics*, 86(1):4–29, 2004.
- G.W. Imbens and D.B. Rubin. *Causal Inference in Statistics, Social, and Biomedical Sciences*. Causal Inference for Statistics, Social, and Biomedical Sciences: An Introduction. Cambridge University Press, 2015. ISBN 9780521885881.
- Toru Kitagawa and Aleksey Tetenov. Who should be treated? empirical welfare maximization methods for treatment choice. *Econometrica*, 86(2):591–616, 2018.
- Paul R Rosenbaum and Donald B Rubin. The central role of the propensity score in observational studies for causal effects. *Biometrika*, 70(1):41–55, 1983.
- Adith Swaminathan and Thorsten Joachims. Batch learning from logged bandit feedback through counterfactual risk minimization. *Journal of Machine Learning Research*, 16:1731–1755, 2015.

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- Baqun Zhang, Anastasios A Tsiatis, Marie Davidian, Min Zhang, and Eric Laber. Estimating optimal treatment regimes from a classification perspective. *Stat*, 1(1):103–114, 2012.
- Yingqi Zhao, Donglin Zeng, A John Rush, and Michael R Kosorok. Estimating individualized treatment rules using outcome weighted learning. *Journal of the American Statistical Association*, 107(499): 1106–1118, 2012.
- Xin Zhou, Nicole Mayer-Hamblett, Umer Khan, and Michael R Kosorok. Residual weighted learning for estimating individualized treatment rules. *Journal of the American Statistical Association*, 112(517): 169–187, 2017.